

STATISTICS 743 (Part I)

The following definitions of common probability distributions, their means and variances may be used without proof in your solutions.

The probability mass or density functions are 0 outside of the values or ranges of values specified below.

- The binomial(n, p) probability mass function is

$$f(x | p) = \binom{n}{x} p^x (1-p)^{n-x} \quad x = 0, 1, \dots, n$$

where n is a positive integer and $0 < p < 1$.

If $X \sim \text{binomial}(n, p)$ then $E(X) = np$ and $\text{Var}(X) = np(1-p)$.

- The Poisson(λ) probability mass function is

$$f(x | \lambda) = \frac{\lambda^x e^{-\lambda}}{x!} \quad x = 0, 1, \dots$$

where $\lambda > 0$

If $X \sim \text{Poisson}(\lambda)$ then $E(X) = \lambda$ and $\text{Var}(X) = \lambda$.

- The Geometric(p) probability mass function is

$$f(x | p) = p(1-p)^{x-1} \quad x = 1, 2, \dots$$

where $0 < p < 1$.

If $X \sim \text{geometric}(p)$ then $E(X) = 1/p$ and $\text{Var}(X) = (1-p)/p^2$.

- The Negative Binomial(r, p) probability mass function is

$$f(x | r, p) = \binom{r+x-1}{x} p^r (1-p)^x \quad x = 0, 1, \dots$$

where r is a positive integer and $0 < p < 1$.

If $X \sim \text{negative binomial}(r, p)$ then $E(X) = r(1-p)/p$ and $\text{Var}(X) = r(1-p)/p^2$.

- The uniform(a, b) probability density function is

$$f(x \mid a, b) = \frac{1}{b - a} \quad a < x < b.$$

where a and b are real numbers with $b > a$.

If $X \sim \text{uniform}(a, b)$ then $E(X) = (a + b)/2$ and $\text{Var}(X) = (b - a)^2/12$.

- The univariate normal(μ, σ^2) probability density function is

$$f(x \mid \mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left\{ -\frac{(x - \mu)^2}{2\sigma^2} \right\} \quad x \in \mathbb{R}$$

where $\mu \in \mathbb{R}$ and $\sigma^2 > 0$.

If $X \sim \text{normal}(\mu, \sigma^2)$ then $E(X) = \mu$ and $\text{Var}(X) = \sigma^2$.

- The exponential(μ) probability density function is

$$f(x; \mu) = \frac{1}{\mu} e^{-x/\mu} \quad x \geq 0.$$

where $\mu > 0$.

If $X \sim \text{exponential}(\mu)$ then $E(X) = \mu$ and $\text{Var}(X) = \mu^2$.

- The gamma(α, β) probability density function is

$$f(x; \alpha, \beta) = \frac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} e^{-x/\beta} \quad x \geq 0.$$

where $\alpha > 0$ and $\beta > 0$ and the gamma function is defined as

$$\Gamma(\alpha) = \int_0^\infty x^{\alpha-1} e^{-x} dx$$

If $X \sim \text{gamma}(\alpha, \beta)$ then $E(X) = \alpha\beta$ and $\text{Var}(X) = \alpha\beta^2$.

- If a random variable X has a gamma distribution with $\beta = 2$ and $\alpha = p/2$ then we say that X has a chi-squared distribution with p degrees of freedom.

- The beta(α, β) probability density function is

$$f(x; \alpha, \beta) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1 - x)^{\beta-1} \quad 0 < x < 1.$$

where $\alpha > 0$ and $\beta > 0$.

If $X \sim \text{beta}(\alpha, \beta)$ then $E(X) = \alpha/(\alpha + \beta)$ and $\text{Var}(X) = \alpha\beta/((\alpha + \beta)^2(\alpha + \beta + 1))$.

The following results may be used without proof in your solutions

- If A and B are any two events then $\Pr(A \cup B) = \Pr(A) + \Pr(B) - \Pr(A \cap B)$.
- If A_1, A_2, \dots is a partition of a sample space and B is any event in the same sample space then

$$P(A_i | B) = \frac{P(B | A_i)P(A_i)}{\sum_{j=1}^{\infty} P(B | A_j)P(A_j)}$$

- If X is a random variable then the moment generating function of X is defined to be $M_X(t) = E(e^{tX})$ provided this expectation exists for t in a neighbourhood of 0.
- $\lim_{n \rightarrow \infty} \left(1 + \frac{a_n}{n}\right)^n = \exp\left\{\lim_{n \rightarrow \infty} a_n\right\}$.
- A distribution is said to be in the exponential family if its probability density or mass function can be written as

$$f(x | \boldsymbol{\theta}) = h(x)c(\boldsymbol{\theta}) \exp\left\{\sum_{i=1}^k w_i(\boldsymbol{\theta})t_i(x)\right\}$$

for positive function $h(x)$ and $c(\boldsymbol{\theta})$.

- A distribution is said to be in a location-scale family if its probability density function can be written as

$$f(x | \mu, \sigma) = \frac{1}{\sigma} g\left(\frac{x - \mu}{\sigma}\right)$$

for some density function g .

- If X and Y are two random variables with joint density (or mass) function $f_{X,Y}$ then the conditional density (or mass) function is defined to be

$$f_{Y|X}(y | X = x) = \frac{f_{X,Y}(x, y)}{f_X(x)}$$

where f_X is the marginal pdf or pmf for X obtained by integrating or summing $f_{X,Y}$ over all possible values of y .

- For any two random variables X and Y we can write $E(X) = E(E(X | Y))$ and $\text{Var}(X) = E(\text{Var}(X | Y)) + \text{Var}(E(X | Y))$.
- The correlation coefficient between two random variables X and Y is

$$\rho_{X,Y} = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X)\text{Var}(Y)}}$$

where the covariance is $\text{Cov}(X, Y) = E((X - E(X))(Y - E(Y))) = E(XY) - E(X)E(Y)$.

- The bivariate normal distribution has joint pdf

$$f(x_1, x_2) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \times \exp \left\{ -\frac{1}{2(1-\rho^2)} \left(\frac{(x_1 - \mu_1)^2}{\sigma_1^2} + \frac{(x_2 - \mu_2)^2}{\sigma_2^2} - 2\rho \frac{(x_1 - \mu_1)(x_2 - \mu_2)}{\sigma_1\sigma_2} \right) \right\}$$

where μ_1 and μ_2 are the marginal means, σ_1^2 and σ_2^2 are the marginal variances and $-1 < \rho < 1$ is the correlation between the two components of the random vector.

- If X_1, \dots, X_n are random variables and a_1, \dots, a_n are constants then

$$\begin{aligned} E \left(\sum_{i=1}^n a_i X_i \right) &= \sum_{i=1}^n a_i E(X_i) \\ \text{Var} \left(\sum_{i=1}^n a_i X_i \right) &= \sum_{i=1}^n a_i^2 \text{Var}(X_i) + 2 \sum_{i>j} a_i a_j \text{Cov}(X_i, X_j) \end{aligned}$$

- $X_n \xrightarrow{d} X$ if $\lim_{n \rightarrow \infty} \Pr(X_n \leq x) = \Pr(X \leq x)$ at every point x at which the cumulative distribution function of X is continuous.
- $X_n \xrightarrow{p} X$ if, for every $\varepsilon > 0$, $\lim_{n \rightarrow \infty} \Pr(|X_n - X| < \varepsilon) = 1$.
- $X_n \xrightarrow{p} X \Rightarrow X_n \xrightarrow{d} X$.
- A sequence X_1, X_2, \dots , is bounded in probability if for every $\varepsilon > 0$ there exist constants B_ε and N_ε such that $n > N_\varepsilon \Rightarrow \Pr(|X_n| < B_\varepsilon) \geq 1 - \varepsilon$.
- If $U \sim \text{Uniform}(0, 1)$ and F is a cumulative distribution function with inverse F^{-1} then $X = F^{-1}(U)$ has cumulative distribution function F .
- A statistic $T(\mathbf{X})$ is sufficient for the parameter θ if the joint pdf (or pmf) of \mathbf{X} can be written as $f(\mathbf{x} | \theta) = g(T(\mathbf{x}), \theta)h(\mathbf{x})$ for every \mathbf{x} .
- The family of sampling distributions of a statistic T is complete if $E_\theta(g(T)) = 0 \Rightarrow \Pr_\theta(g(T) = 0) = 1$ for every possible θ