

## Lecture 5

Ex.

### The Matching Problem Example

This is example 5 in from the textbook.

Idea: Suppose that  $N$  people throw their hats into the centre of the room. The hats are thoroughly mixed and each person picks up a hat. What is the probability that no one gets their own hat?

Event A.

Let's consider some cases:

•  $N=1$ : One person, one hat. He gets his own hat.  $P(A) = 0$ .

•  $N=2$ : two people two hats. The hats can either switch, or go to their respective owners, so  $P(A) = 1/2$ .

So the probability seems to be increasing.

For general  $N$ : we compute  $P(A^c)$

$A^c$  = "at least one person gets their own hat."

Let  $E_i$  be the event:

$E_i$  = the  $i^{\text{th}}$  person gets their own hat.

Therefore,  $A^c = \bigcup_{i=1}^N E_i$  [and  $P(A) = 1 - P(A^c)$ .]

By the inclusion-exclusion principle:

$$\begin{aligned} P\left(\bigcup_{i=1}^N E_i\right) &= \sum_{r=1}^N (-1)^{r+1} \sum_{i_1, \dots, i_r} P(E_{i_1}, \dots, E_{i_r}) \\ &= P(E_1) + \dots + P(E_N) + \dots + (-1)^{N+1} P(E_1, \dots, E_N). \end{aligned}$$

Note that, for each  $r$ , the sum

$$\sum_{i_1, \dots, i_r} P(E_{i_1}, \dots, E_{i_r})$$

has  $\binom{N}{r}$ -many terms.

Let  $i_1, \dots, i_r$  be given/chosen.

$P(E_{i_1}, \dots, E_{i_r})$  is the probability that  $i_1^{\text{th}}, \dots, i_r^{\text{th}}$  people all get their own hats. That means there are  $n-r$  other people to receive hats, and so these hats can be permuted in  $(n-r)!$  many ways. Since there are  $N$  hats total, there are  $N!$  many ways to permute the hats, hence:

$$P(E_1, \dots, E_r) = \frac{(N-r)!}{N!}$$

Thus:

$$\sum_{i_1 < \dots < i_r} P(E_{i_1} \dots E_{i_r}) = \binom{N}{r} \frac{(N-r)!}{N!}$$

$$= \frac{\cancel{N!}}{r! \cancel{(N-r)!}} \cdot \frac{\cancel{(N-r)!}}{\cancel{N!}}$$

$$\text{so } P(A^c) = \sum_{r=1}^N (-1)^{r+1} \frac{1}{r!} = \frac{1}{r!}$$

$$\text{so } P(A) = 1 - \sum_{r=1}^N (-1)^{r+1} \frac{1}{r!}$$

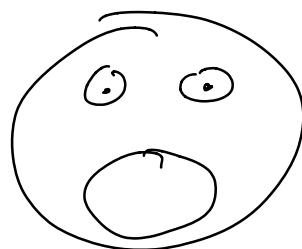
$$= 1 + \sum_{r=1}^N (-1)^r \frac{1}{r!}$$

$$= \sum_{r=0}^N \frac{(-1)^r}{r!}$$

*r, not  
r+1.  
Distributivity  
(-1)*

So as  $N \rightarrow \infty$

$$P(A) \rightarrow \frac{1}{e}.$$



## Probability as a Continuous Set Function

Recall that if  $S$  is a finite sample space, then for  $E \subseteq S$

$$P(E) = \sum_{x \in E} P(\{x\}).$$

and so, if  $F \subseteq E$  and  $P(F) = P(E)$ , then  $F = E$ .

This sort of reasoning quickly breaks down, especially in the uncountable case:

Imagine the experiment where you randomly select a real number from  $[0, 1]$ . What is the probability  $P(x \in [0, \frac{1}{2}])$ ? Since  $[0, \frac{1}{2}]$  is "half" of the sample space, our intuition tells us that  $P(x \in [0, \frac{1}{2}])$  should be  $\frac{1}{2}$ .

But what about  $P(x \in (0, \frac{1}{2}))$ ? The set  $(0, \frac{1}{2})$  is a proper subset of  $[0, \frac{1}{2}]$ . So the probability "should" be less... but it isn't.

Question: what is  $P(x=0)$ ?

Defn: We say that a sequence of events  $\{E_n : n \geq 0\}$  is increasing if  $E_0 \subseteq E_1 \subseteq \dots \subseteq E_n \subseteq \dots$  (strictly increasing if  $E_i \subsetneq E_{i+1} \forall i$ ).

and decreasing if  $E_0 \supseteq E_1 \supseteq \dots \supseteq E_n \supseteq \dots$

If  $\{E_n\}$  is an increasing sequence, we define

$$\lim_{n \rightarrow \infty} E_n = \bigcup_{n=0}^{\infty} E_n$$

If  $\{E_n\}$  is decreasing, we say

$$\lim_{n \rightarrow \infty} E_n = \bigcap_{n=0}^{\infty} E_n$$

Proposition If  $\{E_n\}$  is increasing or

decreasing then

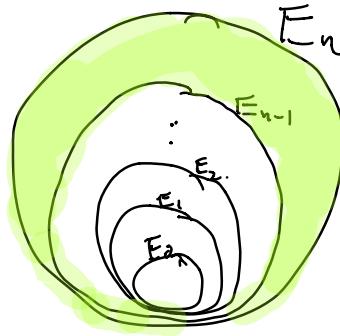
$$P\left(\lim_n E_n\right) = \lim_n P(E_n).$$

Pf: We will prove it in the case of increasing.

For each  $n$ , let

$$F_n = E_n \left( \bigcup_{i=0}^{n-1} E_i \right)^c$$

$$F_0 = E_0.$$



So

$$\bigcup_{i=0}^n F_i = \bigcup_{i=0}^n E_n = \underbrace{E_n}_{\text{since increasing.}}$$

and hence  $\bigcup_{i=0}^{\infty} F_i = \bigcup_{i=0}^{\infty} E_i$ .

But the  $F_i$  are mutually exclusive,  
so by the axioms:

$$\begin{aligned} P\left(\bigcup_{i=0}^{\infty} E_i\right) &= P\left(\bigcup_{i=0}^{\infty} F_i\right) \\ &= \sum_{i=0}^{\infty} P(F_i) \quad (\text{by the last axiom}) \\ &= \lim_{n \rightarrow \infty} \sum_{i=0}^n P(F_i) \\ &= \lim_{n \rightarrow \infty} P\left(\bigcup_{i=1}^n F_i\right) \\ &= \lim_{n \rightarrow \infty} P(E_n). \end{aligned}$$




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Now, consider the following increasing sequence

$$E_i = [\frac{1}{i}, 1] \subseteq [0, 1]. \quad (i=1, 2, \dots)$$

In the game of choosing a random # in  $[0, 1]$

$$P(x \in E_i) \geq 1 - \frac{1}{i}$$

$$\text{So } \lim_{n \rightarrow \infty} P(x \in E_n) = \lim_{n \rightarrow \infty} 1 - \frac{1}{n} = 1.$$

$$\text{But } E = \lim_{n \rightarrow \infty} E_n = \bigcup_{n=1}^{\infty} E_n = (0, 1].$$

$$\text{So } P(E^c) = 1 - 1 = 0.$$

$$\text{and } E^c = \{x \in [0, 1] : x \notin (0, 1]\}$$

$$\text{i.e. } E^c = \{0\}.$$

$$\text{Hence } P(x=0) = 0.$$

What about for any real  $r \in [0, 1]$ ?

$$\text{what } \geq P(x=r) ?$$