

## Lecture 23

### Last time:

- We want to find a confidence interval on the mean  $\mu$  in the situation where we are sampling a normal distribution with unknown variance.

- In this situation, we consider

$$T = \frac{\bar{X} - \mu}{S/\sqrt{n}},$$

the student's t-distribution with  $(n-1)$  degrees of freedom.

- Here,  $S$  is the sample standard deviation,  $n$  is the sample size.
- if  $\bar{x}$  is a sample mean, then a  $100(1-\alpha)\%$  CI for  $\mu$  is given by

$$P\left(\bar{x} - t_{\alpha/2, n-1} \frac{S}{\sqrt{n}} \leq \mu \leq \bar{x} + t_{\alpha/2, n-1} \frac{S}{\sqrt{n}}\right) = 1-\alpha.$$

where  $t_{\alpha/2, n-1}$  is the value of  $T = \frac{\bar{X} - \mu}{S/\sqrt{n}}$  such that  $P(T \leq t_{\alpha/2, n-1}) = 1 - \alpha/2$ .

Exx: Out of all the Stats 31/3J students 12 wrote their upcoming test at an alternate time. For these students, we found that  $\bar{x} = 77\%$  with  $S = 13.6\%$ . Find a 95% CI for the average of the whole class.

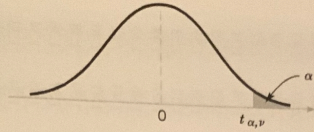
Soln: First, we find  $\alpha$ :  $100(1-\alpha) = 95$   
 $\Downarrow$   
 $1-\alpha = 0.95 \rightarrow \alpha = 0.05.$

We have that the endpoints of our interval are  $\bar{x} \pm \left(t_{\alpha/2, \frac{12-1}{2}}\right) \frac{S}{\sqrt{n}} = 77 \pm \left(t_{0.025, 11}\right) \frac{13.6}{\sqrt{12}}$

So we need only to find  $t_{0.025, 11}$ :

From the chart,  $t_{0.025, 11} = 2.201.$

Hence the CI is  $77 \pm (2.201) \frac{13.6}{\sqrt{12}}$   
 $= 77 \pm \underbrace{8.647}_{\text{Error.}}$

TABLE V Percentage Points  $t_{\alpha, \nu}$  of the  $t$  Distribution

| $\alpha$<br>$\nu$ | .40  | .25   | .10   | .05   | .025   | .01    | .005   | .0025  | .001   | .0005  |
|-------------------|------|-------|-------|-------|--------|--------|--------|--------|--------|--------|
| 1                 | .325 | 1.000 | 3.078 | 6.314 | 12.706 | 31.821 | 63.657 | 127.32 | 318.31 | 636.62 |
| 2                 | .289 | .816  | 1.886 | 2.920 | 4.303  | 6.965  | 9.925  | 14.089 | 23.326 | 31.598 |
| 3                 | .277 | .765  | 1.638 | 2.353 | 3.182  | 4.541  | 5.841  | 7.453  | 10.213 | 12.924 |
| 4                 | .271 | .741  | 1.533 | 2.132 | 2.776  | 3.747  | 4.604  | 5.598  | 7.173  | 8.610  |
| 5                 | .267 | .727  | 1.476 | 2.015 | 2.571  | 3.365  | 4.032  | 4.773  | 5.893  | 6.869  |
| 6                 | .265 | .718  | 1.440 | 1.943 | 2.447  | 3.143  | 3.707  | 4.317  | 5.208  | 5.959  |
| 7                 | .263 | .711  | 1.415 | 1.895 | 2.365  | 2.998  | 3.499  | 4.029  | 4.785  | 5.408  |
| 8                 | .262 | .706  | 1.397 | 1.860 | 2.306  | 2.896  | 3.355  | 3.833  | 4.501  | 5.041  |
| 9                 | .261 | .703  | 1.383 | 1.833 | 2.262  | 2.821  | 3.250  | 3.690  | 4.297  | 4.781  |
| 10                | .260 | .700  | 1.372 | 1.812 | 2.228  | 2.764  | 3.169  | 3.581  | 4.144  | 4.587  |
| 11                | .260 | .697  | 1.363 | 1.796 | 2.201  | 2.718  | 3.106  | 3.497  | 4.025  | 4.437  |
| 12                | .259 | .695  | 1.356 | 1.782 | 2.179  | 2.681  | 3.055  | 3.428  | 3.930  | 4.318  |
| 13                | .259 | .694  | 1.350 | 1.771 | 2.160  | 2.650  | 3.012  | 3.372  | 3.852  | 4.221  |
| 14                | .258 | .692  | 1.345 | 1.761 | 2.145  | 2.624  | 2.977  | 3.326  | 3.787  | 4.140  |
| 15                | .258 | .691  | 1.341 | 1.753 | 2.131  | 2.602  | 2.947  | 3.286  | 3.733  | 4.073  |
| 16                | .258 | .690  | 1.337 | 1.746 | 2.120  | 2.583  | 2.921  | 3.252  | 3.686  | 4.015  |
| 17                | .257 | .689  | 1.333 | 1.740 | 2.110  | 2.567  | 2.898  | 3.222  | 3.646  | 3.965  |
| 18                | .257 | .688  | 1.330 | 1.734 | 2.101  | 2.552  | 2.878  | 3.197  | 3.610  | 3.922  |
| 19                | .257 | .688  | 1.328 | 1.729 | 2.093  | 2.539  | 2.861  | 3.174  | 3.579  | 3.883  |
| 20                | .257 | .687  | 1.325 | 1.725 | 2.086  | 2.528  | 2.845  | 3.153  | 3.552  | 3.850  |
| 21                | .257 | .686  | 1.323 | 1.721 | 2.080  | 2.518  | 2.831  | 3.135  | 3.527  | 3.819  |
| 22                | .256 | .686  | 1.321 | 1.717 | 2.074  | 2.508  | 2.819  | 3.119  | 3.505  | 3.792  |
| 23                | .256 | .685  | 1.319 | 1.714 | 2.069  | 2.500  | 2.807  | 3.104  | 3.485  | 3.767  |
| 24                | .256 | .685  | 1.318 | 1.711 | 2.064  | 2.492  | 2.797  | 3.091  | 3.467  | 3.745  |
| 25                | .256 | .684  | 1.316 | 1.708 | 2.060  | 2.485  | 2.787  | 3.078  | 3.450  | 3.725  |
| 26                | .256 | .684  | 1.315 | 1.706 | 2.056  | 2.479  | 2.779  | 3.067  | 3.435  | 3.707  |
| 27                | .256 | .684  | 1.314 | 1.703 | 2.052  | 2.473  | 2.771  | 3.057  | 3.421  | 3.690  |
| 28                | .256 | .683  | 1.313 | 1.701 | 2.048  | 2.467  | 2.763  | 3.047  | 3.408  | 3.674  |
| 29                | .256 | .683  | 1.311 | 1.699 | 2.045  | 2.462  | 2.756  | 3.038  | 3.396  | 3.659  |
| 30                | .256 | .683  | 1.310 | 1.697 | 2.042  | 2.457  | 2.750  | 3.030  | 3.385  | 3.646  |
| 40                | .255 | .681  | 1.303 | 1.684 | 2.021  | 2.423  | 2.704  | 2.971  | 3.307  | 3.551  |
| 60                | .254 | .679  | 1.296 | 1.671 | 2.000  | 2.390  | 2.660  | 2.915  | 3.232  | 3.460  |
| 120               | .254 | .677  | 1.289 | 1.658 | 1.980  | 2.358  | 2.617  | 2.860  | 3.160  | 3.373  |
| $\infty$          | .253 | .674  | 1.282 | 1.645 | 1.960  | 2.326  | 2.576  | 2.807  | 3.090  | 3.291  |

 $\nu$  = degrees of freedom.

## Confidence Intervals for Population Proportion.

- Suppose we take a population and we are interested in a certain subclass
  - ↳ e.g. the population could be 31/3J students and the subclass of interest could be the collection of students dressing up for Halloween.
- Let  $p$  be the proportion of the population in the class of interest.
- Suppose we take a large sample  $n$ . Let  $X (\leq n)$  be the number of samples which fall into this class.
- Then  $\hat{P} = \frac{X}{n}$  is a point estimator of  $p$
- Note that  $X$  is binomial  $Bin(n, p)$  and so  $E[\hat{P}] = E\left[\frac{X}{n}\right] = \frac{1}{n} E[X] = \frac{1}{n} (np) = p$ .
  - ↳ i.e.  $\hat{P}$  is an unbiased estimator of  $p$ .
- For large  $n$ , we have that
$$Z = \frac{X - np}{\sqrt{np(1-p)}} = \frac{\hat{P} - p}{\sqrt{\frac{p(1-p)}{n}}}$$
is approximately the standard normal.
- Suppose now that we want to find a  $100(1-\alpha)\%$  CI for  $p$ .

- taking  $z_{\alpha/2}$  s.t.  $P(Z \leq z_{\alpha/2}) = 1 - \alpha/2$ , we have that

$$P(-z_{\alpha/2} \leq \frac{\hat{P} - p}{\sqrt{\frac{p(1-p)}{n}}} \leq z_{\alpha/2}) \approx 1 - \alpha$$

and so

$$P(\hat{P} - z_{\alpha/2} \sqrt{\frac{p(1-p)}{n}} \leq p \leq \hat{P} + z_{\alpha/2} \sqrt{\frac{p(1-p)}{n}}) \approx 1 - \alpha.$$

- Here,  $\sqrt{\frac{p(1-p)}{n}}$  is the standard error of  $\hat{P}$ ,  $\sigma_{\hat{P}}$  (from chapter 7).

- This is not ideal, since the upper and lower confidence bounds contain  $p$ , which is unknown.

↳ however we can approximate  $p$  by  $\hat{P}$ , giving:

$$P(\hat{P} - z_{\alpha/2} \sqrt{\frac{\hat{P}(1-\hat{P})}{n}} \leq p \leq \hat{P} + z_{\alpha/2} \sqrt{\frac{\hat{P}(1-\hat{P})}{n}}) \approx 1 - \alpha.$$

approx.  $100(1-\alpha)\%$  CI on  $p$ .

- Note that this procedure depends on having a good approximation to the binomial distribution, and so we want  $np \geq 5$  and  $n(1-p) \geq 5$ .

Ex: Out of a random sample of 50 engineers



Students, 18 are dressing up for Halloween.  
Find a 99% confidence interval for the proportion of eng students who are dressing up for Halloween.

Soln:  $\alpha = 0.01$ , So  $1 - \alpha/2 = 0.995$ .

Z-scores:  $Z_{\alpha/2} = 2.58$ .

Now the point estimate  $\hat{p}$  for  $P$  is  
 $\hat{p} = 18/50 = 0.36$ .

So the 99% CI is  $\hat{p} \pm Z_{\alpha/2} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$   
 $= 0.36 \pm (2.58) \sqrt{\frac{0.36(1-0.36)}{50}}$   
 $= 0.36 \pm 0.175$

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Now, suppose we want a  $100(1-\alpha)\%$  CI for the proportion with a certain amount of error,  $E = Z_{\alpha/2} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$ .

- what should  $n$  be?

$$n = \left[ \frac{Z_{\alpha/2}^2}{E^2} (\hat{p}(1-\hat{p})) \right]$$

but what if we don't know  $\hat{p}$ ??

2 methods:

① Use an earlier study and compute  $\hat{p}$  from that (eg: in the last example, found  $\hat{p} = 0.36$ ; use that).

② Observe that  $p(1-p) \leq 0.25$  (when  $p=0.5$ ) and so substitute  $\hat{p} = 0.5$  as a worst case scenario:  $n = \left\lceil \left( \frac{z_{\alpha/2}}{E} \right)^2 (0.25) \right\rceil$

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Ex: We want a 95% CI for the proportion of eng students dressing up for halloween with error at most  $E=0.07$ . What should the sample size be?

$$\alpha = 0.05 \rightsquigarrow z_{\alpha/2} = 1.96.$$

$$\text{so } n = \left\lceil \left( \frac{1.96}{0.07} \right)^2 (\hat{p}(1-\hat{p})) \right\rceil.$$

method ①: In the small test study (sample size 50) we had  $\hat{p} = 0.36$ .

$$\text{so } n \geq \left\lceil \left( \frac{1.96}{0.07} \right)^2 (0.36)(1-0.36) \right\rceil = \left\lceil 180.36 \right\rceil = 181$$

method ②: (worst case scenario)  $\hat{p}(1-\hat{p}) = 0.25$

$$\text{so } n \geq \left\lceil \left( \frac{1.96}{0.07} \right)^2 (0.25) \right\rceil = 196.$$