

## Lecture 28

Tests on the mean of  
a normal distr. w/ known  
variance:  $Z_0 = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}}$

### Large Sample Sizes

- By the central limit theorem, if  $n$  is large (usually  $n \geq 30$ ) then we can drop the assumption that we are sampling a normal distribution (since  $\bar{X}$  will be approximately normal anyway).
- If the variance is unknown, but the sample size is large enough, say  $n \geq 40$ , then we may substitute  $s^2$  (sample variance) for  $\sigma^2$  in all of the above without changing much.  
*really s*

### Hypothesis Testing on the mean of a normal distribution, unknown variance, small sample size.

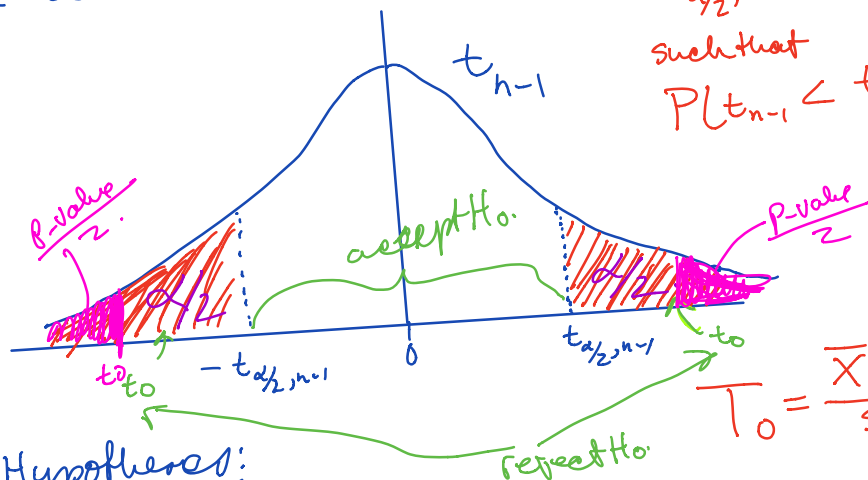
- In this situation we take an approach analogous to what we did for confidence intervals.
- Consider a null hypothesis  $H_0: \mu = \mu_0$ .
- we choose as a test statistic

$$T_0 = \frac{\bar{X} - \mu_0}{s/\sqrt{n}}$$

where  $n$  is the sample size, and  $s$  is the sample standard deviation.

- If  $H_0$  is true, then  $T_0 = t_{n-1}$ , i.e.  $T_0$  has a Student's  $t$ -distribution with  $n-1$  degrees of freedom (so these tests are called "t-tests").

- choose  $\alpha$  a significance level.



$t_{\alpha/2, n-1}$  is the value such that  $P(t_{n-1} < t_{\alpha/2, n-1}) = 1 - \frac{\alpha}{2}$ .

$$T_0 = \frac{\bar{X} - \mu_0}{s/\sqrt{n}}$$

Alternate Hypotheses:

①  $H_1: \mu \neq \mu_0$  (two-sided).

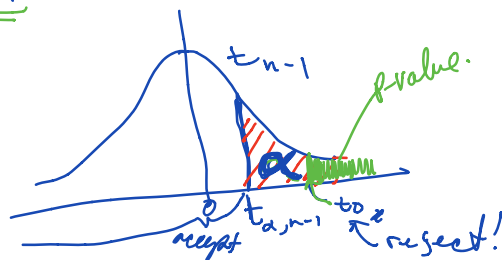
Critical region:  $|t_0| > t_{\alpha/2, n-1}$  ← reject  $H_0$ !

P-value:  $2P(T_0 > |t_0|)$

②  $H_1: \mu > \mu_0$

Critical region:  $t_0 > t_{\alpha, n-1}$  ← reject  $H_0$ !

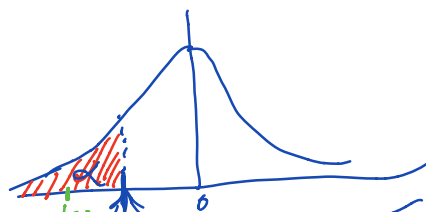
P-value:  $P(T_0 > t_0)$



③  $H_1: \mu < \mu_0$

Critical region:  $t_0 < -t_{\alpha, n-1}$

P-value:  $P(T_0 < t_0)$



reject if  $t_0$  is here. accept if  $t_0$  is here  
 $-t_{\alpha, n-1}$

One thing to note!

↳ when performing  $t$ -tests, statistical software can give  $P$ -values precisely.

↳ when doing  $t$ -tests by hand, this is almost impossible, since the  $t$ -table are too coarse.

Ex: consider the following situations:

a)  $H_0: \mu = \mu_0$

$H_1: \mu \neq \mu_0$

$t_0 = 2.537$

$n = 10$

b)  $H_0: \mu = \mu_0$

$H_1: \mu > \mu_0$

$t_0 = 1.863$

$n = 16$

Let's estimate the  $P$ -values in these two situations:

a)  $n = 10, \rightarrow 10 - 1 = 9$  degrees of freedom: (2-sided case).

TABLE V Percentage Points  $t_{\alpha, \nu}$  of the  $t$  Distribution

$\alpha$	.40	.25	.10	.05	.025	.01	.005	.0025	.001	.0005
1	.325	1.000	3.078	6.314	12.706	31.821	63.657	127.32	318.31	636.62
9	.261	.703	1.383	1.833	2.262	2.821	3.250	3.690	4.297	4.781

degrees of freedom

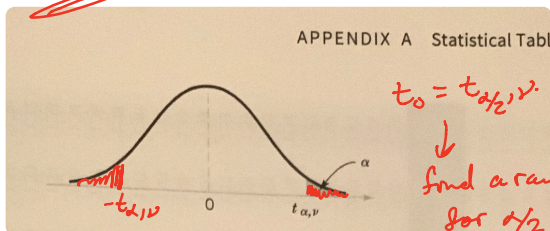
$t_0 = 2.537$

$\Rightarrow P\text{-value}$

is between 0.01 and 0.025

So  $0.02 \leq P\text{-value} \leq 0.05$

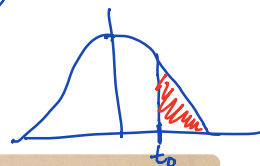
i.e.  $0.02 \leq P\text{-value} \leq 0.05$   
 $\underbrace{\quad}_{2(0.01)} \quad \underbrace{\quad}_{2(0.025)}$



b)  $n = 16$ , so 15 degrees of freedom.

$t_0 = 1.863$  (1-sided case)

$H_1: \mu > \mu_0$



*p-value in this range.*  
 $\alpha = p\text{-value}$

$\alpha \backslash v$	.40	.25	.10	.05	.025	.01	.005	.0025	.001	.0005
1	.325	1.000	3.078	6.314	12.706	31.821	63.657	127.32	318.31	636.62
15	.258	.691	1.341	1.753	2.131	2.602	2.947	3.286	3.733	4.073

*degrees of freedom* (pointing to 15)  
 $t_0 = 1.863$  in this range.

So  $p\text{-value}$  is between 0.025 and 0.05.

## Tests on Population proportion

- Suppose some population has some particular subclass which is proportion  $p$  of the total population:

↳ eg: of all engineering students what proportion are female?

- Suppose we take a sample of size  $n$ , and let  $X = \#$  of samples from the subclass.

- Then  $X \sim \text{Bin}(n, p)$  and  $\hat{p} = \frac{X}{n}$  is an unbiased estimator of  $p$  (See notes on CI's for pop. proportion).



- If  $np > 5$ ,  $n(1-p) > 5$ , then we may further approximate  $X \approx N(np, np(1-p))$ .

Suppose our test is:

$$H_0: p = p_0$$

$$H_1: \textcircled{I} p \neq p_0, \textcircled{II} p > p_0, \textcircled{III} p < p_0.$$

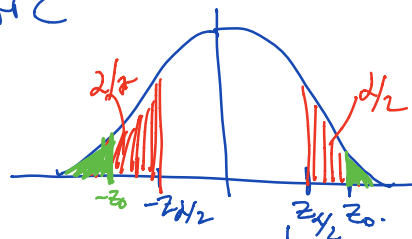
Significance level  $\alpha$ .

Sub.  $p = p_0$

- If  $H_0$  is true, then  $X \approx N(np_0, np_0(1-p_0))$   
and so  $\hat{p} = \frac{X}{n} \approx N(p_0, \frac{p_0(1-p_0)}{n})$

- We choose our test statistic

$$Z_0 = \frac{\hat{p} - p_0}{\sqrt{\frac{p_0(1-p_0)}{n}}}$$



- Then  $Z_0 \approx N(0, 1)$  iff  $H_0$  is true!

- Suppose, for some sample,  $\underline{Z_0 = z_0}$ . Then:

$\textcircled{I} H_1: p \neq p_0$   
(two-sided).

P-value:  $\frac{2P(Z_0 \geq |z_0|)}{= 2\Phi(-|z_0|)}$

Reject  $H_0$  if  $|z_0| \geq z_{\alpha/2}$ .

$\textcircled{II} H_1: p > p_0$

P-value:  $\frac{P(Z_0 \geq z_0)}{= \Phi(-z_0)}$

Reject  $H_0$  iff  $z_0 \geq z_{\alpha}$ .

$\textcircled{III} H_1: p < p_0$

P-value:  $P(\underline{Z_0} < z_0) = \Phi(z_0)$

Reject  $H_0$  iff  $z_0 \leq -z_{\alpha}$ .

Example: In 2018, McMaster claimed that 27% of all incoming engineering students were female. Of a random sample of 25 engineering students, 4 are female. At a significance level of  $\alpha = 0.01$ , does the evidence support McMaster's claim?

Solution:  $\hat{p} = \frac{4}{25} = 0.16$

$H_0: p = 0.27$

$H_1: p < 0.27$

$$Z_0 = \frac{\hat{p} - p_0}{\sqrt{\frac{p_0(1-p_0)}{n}}} = \frac{0.16 - 0.27}{\sqrt{\frac{0.27(0.73)}{25}}} = -1.2389$$

$\hat{p}$        $p_0$   
 $\nwarrow$        $\nwarrow$   
 $p_0$        $1-p_0$

For  $\alpha = 0.01$ ,  $Z_\alpha = 2.33$

so  $-Z_\alpha = -2.33$

$Z_0 \geq -Z_\alpha$   
 $-1.2389 \geq -2.33$

So don't reject  $H_0$ . ✓

$Z_\alpha$  is the value such that  $P(Z < Z_\alpha) = 1 - \alpha$  from Z-chart.