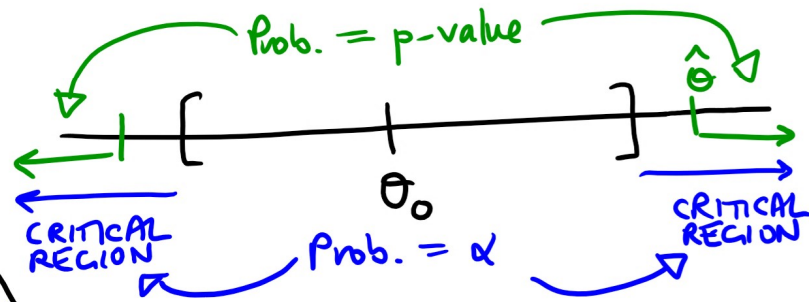


3703-3J04 PROBABILITY & STATISTICS FOR CO1 - Lecture 27 (CIVIL) ENGINEERING

Last time

P-VALUE



TEST ON MEAN OF NORMAL, VARIANCE KNOWN $N(\mu, \sigma^2)$

Reject H_0 if $p\text{-value} < \alpha$.

$$H_0: \mu = \mu_0$$

$$H_1: \text{(I) } \mu \neq \mu_0 \text{ or (II) } \mu > \mu_0 \text{ or (III) } \mu < \mu_0$$

Use test stat. $Z_0 = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}}$

Case I p-value: $2 \cdot P(Z_0 > \frac{|\bar{X} - \mu_0|}{\sigma/\sqrt{n}})$

Reject H_0 if $\frac{|\bar{X} - \mu_0|}{\sigma/\sqrt{n}} > z_{\alpha/2}$ or...

$-\frac{|\bar{X} - \mu_0|}{\sigma/\sqrt{n}} < -z_{\alpha/2}$

if $|z_0| > z_{\alpha/2}$ or

$-|z_0| < -z_{\alpha/2}$

Case II

Reject H_0 for H_1

when $z_0 > z_\alpha$

Case III

Reject H_0 for H_1 , when $z_0 < -z_\alpha$.

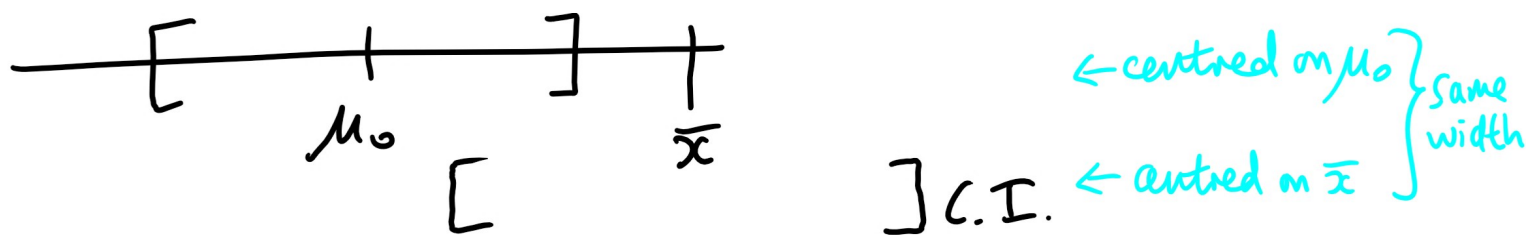
Since reference is $N(0,1)$ the above is called a z-test.

If we unwrap above, we get ^{that} critical region(s) closely resembles the $100(1-\alpha)\%$ C.I. bounds but relative to μ_0 not \bar{x} :

$$\text{I) } H_1 : \mu \neq \mu_0 : CR \bar{x} < \mu_0 - z_{\alpha/2} \left(\frac{\sigma}{\sqrt{n}} \right) \\ \text{or } \bar{x} > \mu_0 + z_{\alpha/2} \left(\frac{\sigma}{\sqrt{n}} \right)$$

$$\text{II) } H_1 : \mu > \mu_0 : CR \bar{x} > \mu_0 + z_{\alpha} \left(\frac{\sigma}{\sqrt{n}} \right)$$

$$\text{III) } H_1 : \mu < \mu_0 : CR \bar{x} < \mu_0 - z_{\alpha} \left(\frac{\sigma}{\sqrt{n}} \right)$$



Example A sample of 25 adults gave mean body temp. of 36.8°C . Body temp. assumed to be Normal with $\sigma = 0.34^{\circ}\text{C}$.

Test $H_0 : \mu = 37^{\circ}\text{C}$

$H_1 : \mu \neq 37^{\circ}\text{C}$ at level $\alpha = 0.01$.

Solution $z_0 = \frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}} = \frac{36.8 - 37}{0.34/5} = -2.94$

P-value : $P(Z_0 \leq -2.94) + P(Z_0 \geq 2.94)$

$= 2(0.001641) = 0.003282$

$< \alpha = 0.01$.

In this case, reject H_0 (mean is unlikely to be 37°C).

Recall $\beta = P(\text{Type II error})$

We fix $\alpha = P(\text{Type I error})$ & this affects β
(α & β in opposition)
but we can still influence β by changing sample size n .

β depends on true mean $\mu = \mu_0 + \delta$, say.
(so $\mu_0 = \mu - \delta$)

So how does β depend on δ .

In 2-sided case:

$$\beta(\delta) = P\left(-z_{\frac{\alpha}{2}} < Z_0 < z_{\frac{\alpha}{2}}\right)$$

$$\text{When } Z_0 = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} = \frac{\bar{X} - (\mu - \delta)}{\sigma/\sqrt{n}}$$

$$= \underbrace{\frac{\bar{X} - \mu}{\sigma/\sqrt{n}}}_{\sim N(0,1)} + \frac{\delta}{\sigma/\sqrt{n}}$$

$$\sim N\left(\frac{\delta}{\sigma/\sqrt{n}}, 1\right)$$

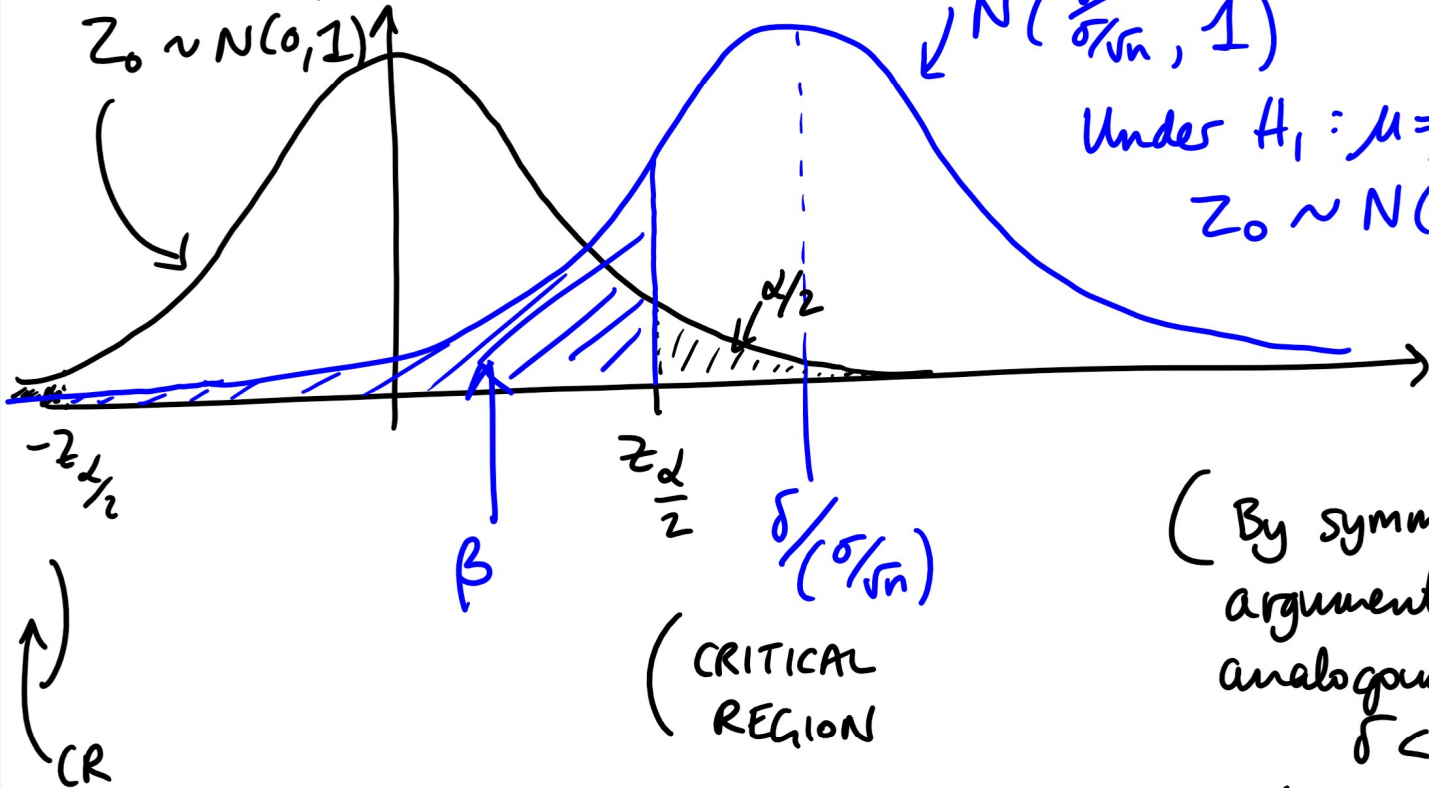
$$\beta(\delta) = P\left(-z_{\alpha/2} - \frac{\delta}{\sigma/\sqrt{n}} < Z < z_{\alpha/2} - \frac{\delta}{\sigma/\sqrt{n}}\right)$$

(standardizing)

$$= \Phi\left(z_{\alpha/2} - \frac{\delta}{(\sigma/\sqrt{n})}\right) - \underbrace{\Phi\left(\frac{-z_{\alpha/2} - \delta}{(\sigma/\sqrt{n})}\right)}_{\approx 0}$$

Under $H_0: \mu = \mu_0$

$$Z_0 \sim N(0, 1)$$



Under $H_1: \mu = \mu_0 + \delta$

$$Z_0 \sim N\left(\frac{\delta}{(\sigma/\sqrt{n})}, 1\right)$$

(By symmetry argument works analogously for $\delta < 0$)

& get an analogue for 1-sided situation (replace $z_{\alpha/2}$ with z_{α} as appropriate)

From above we have

$$z_{\alpha/2} - \frac{\delta}{\sigma/\sqrt{n}} \approx -z_{\beta}$$

So if we fix β & we know $\mu = \mu_0 + \delta$

we find n :

$$\left(\frac{z_{\alpha/2} + z_{\beta}}{\delta}\right)^2 \sigma^2 \approx n$$

Example Battery life in hours $\sim N(\mu, 1.5)$

10 batteries have sample mean 40.5 hours. If the true mean life is 43 hours, how big a sample do

We need to ensure $\beta < 0.05$ in a 2-sided test with $H_0: \mu = 40$, $H_1: \mu \neq 40$ at $\alpha = 0.025$?

Solution

$$n \approx \left(\frac{z_{\alpha/2} + z_{\beta}}{\delta} \right)^2 \sigma^2$$

$$\begin{array}{cc} \mu & \mu_0 \\ | & | \\ \delta & = 43 - 40 \\ & = 3 \end{array}$$

$$= \left(\frac{z_{0.0125} + z_{0.05}}{3} \right)^2 (1.5)$$

For $\left. \begin{array}{l} z_{0.0125} = 2.24 \\ z_{0.05} = 1.64 \end{array} \right\} \rightarrow = \left(\frac{2.24 + 1.64}{3} \right)^2 (1.5) = 2.51$

So $n = 3$ enough.

9.3 Hypothesis Test on Normal Mean, Variance

Unknown

Analogous to C.I. setting —

Use t -distribution.

Here analogous test statistic is

$$T_0 = \frac{\bar{X} - \mu_0}{S / \sqrt{n}}$$

which has a t -distribution with $n-1$ degrees of freedom

All setup completely analogous to known variance

setting: ← but with t_{n-1} distribution instead of $N(0,1)$ distribution.

$$\frac{\bar{x} - \mu_0}{S/\sqrt{n}}$$

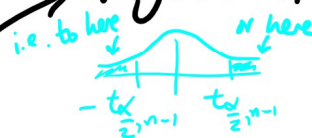
$$H_0 : \mu = \mu_0$$

$$H_1 : \text{(I) } \mu \neq \mu_0$$

$$\text{(II) } \mu > \mu_0$$

$$\text{(III) } \mu < \mu_0$$

Reject H_0 if $|t_0| > t_{\frac{\alpha}{2}, n-1}$ or $-|t_0| < -t_{\frac{\alpha}{2}, n-1}$



Reject H_0 if $t_0 > t_{\alpha, n-1}$



Reject H_0 if $t_0 < -t_{\alpha, n-1}$

