

3703-3704 PROBABILITY & STATISTICS FOR
CO1 - Lecture 38 (CIVIL) ENGINEERING

Today REVIEW

- One last thought on bias.
 - Confidence intervals/tests → when to use which one?
 - Normal Approximation to Binomial
→ don't forget the continuity correction!
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Example Suppose $\left(\frac{X}{n}\right)^2 = \hat{p}^2$ is used to estimate p^2 , where p = proportion of underlying population in a certain category.

Find the bias in \hat{p}^2 .

Solution We need to find $E(\hat{p}^2) - p^2 = \text{bias}$.

$$E\left(\left(\frac{X}{n}\right)^2\right) = E\left(\frac{X^2}{n^2}\right) = \frac{1}{n^2} E(X^2)$$

↑ linear function of r.v. X^2

Cannot use our "linear functions of r.v.s" formulas to find $E(X^2)$.

$$V(X) = E((X - \mu)^2) = E(X^2) - (E(X))^2$$

\uparrow
 $E(X)$

So for any X : $E(X^2) = V(X) + (E(X))^2$

Here $X \sim \text{Bin}(n, p)$

So $E(X) = np$

$V(X) = np(1-p)$

So

$E(X^2)$

$= np(1-p) + n^2 p^2$

Hence $E\left(\frac{X^2}{n^2}\right) = \frac{1}{n^2} E(X^2) = \frac{1}{n^2} (np(1-p) + n^2 p^2)$

\uparrow

$= \frac{p(1-p)}{n} + p^2$

Bias = $E\left(\left(\frac{X}{n}\right)^2\right) - p^2 = \frac{p(1-p)}{n} + \cancel{p^2} - \cancel{p^2}$

Confidence Intervals & Test Statistics

(How to navigate lines 20-29 of Formula Sheet.)

In pairs

C.I.

Test

20) z - C.I. for mean μ

23) z - test for μ

21) t - C.I. for μ

24) t - test for μ

22) (z-) C.I. for proportion p

25) z - test for p

$\left. \begin{matrix} 26) \\ 27) \end{matrix} \right\} (t-) \text{ C.I. for difference of 2 means } \mu_1 - \mu_2 \text{ with variance unknown}$
 $\left. \begin{matrix} 28) \\ 29) \end{matrix} \right\} t\text{-test for } \mu_1 - \mu_2 \text{ with var. unknown.}$

First check what your C.I. / test is for:
 mean μ , proportion p , difference in 2 means $\mu_1 - \mu_2$

① C.I.s / Tests for mean μ

z-tests / C.I.s

(A) Underlying pop: Normal
 Variance σ^2 : Known
 Sample size n : irrelevant

C.I.: $\bar{x} \pm z_{\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}}$ (20)

Test: $z_0 = \frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}}$ (23)

 ← assumed in H_0

t-tests / C.I.s

(B) Underlying pop: Normal
 Variance σ^2 : UNknown
 Sample size n : small

($n < 40$)

(21) C.I.: $\bar{x} \pm t_{\frac{\alpha}{2}, n-1} \frac{s}{\sqrt{n}}$

(24) Test: $t_0 = \frac{\bar{x} - \mu_0}{s/\sqrt{n}}$

(C) Underlying pop: ~~Normal~~ irrelevant (because use C.L.T.)
 Variance σ^2 : UNknown (est. with s^2) (S^2 used to estimate σ^2)
 Sample size n : large ($n > 40$)

C.I. $\bar{x} \pm z_{\frac{\alpha}{2}} \frac{s}{\sqrt{n}}$ H_0 -value

(Test: $z_0 = (\bar{x} - \mu_0) / (s/\sqrt{n})$)

← Not on formula sheet
 — recover from l. 20 with s in place of σ

② C.I.s / Tests for proportion

→ No choice! z-C.I. (22) $\hat{p} \pm z_{\frac{\alpha}{2}} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$

z-test (25) $z_0 = \frac{\hat{p} - p_0}{\sqrt{\frac{p_0(1-p_0)}{n}}}$ ← under p_0

③ C.I.s / Tests on difference in means $\mu_1 - \mu_2$ v. 0
when variance σ^2 is unknown

→ Always a t-test

Question: Are we allowed to assume $\sigma_1^2 = \sigma_2^2$
 ("variances equal")

or not allowed to assume $\sigma_1^2 = \sigma_2^2$ so
 have to assume $\sigma_1^2 \neq \sigma_2^2$ ("variances unequal")?

Equal: (26.) C.I. : $\bar{x}_1 - \bar{x}_2 \pm t_{\frac{\alpha}{2}, n_1+n_2-2} S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}$

(28.) Test: $t_{n_1+n_2-2} = \frac{\bar{x}_1 - \bar{x}_2}{S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$

↑ $S_p^2 =$ pooled sample variance (26.)

Unequal: (27.) C.I. $\bar{x}_1 - \bar{x}_2 \pm t_{\frac{\alpha}{2}, \nu} \sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}$

(28.) Test: $t_{\nu} = \frac{\bar{x}_1 - \bar{x}_2}{\sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}}$

↑ degrees of freedom (27.)

Critical Values & P-Values

$z_{\frac{\alpha}{2}}$ or $t_{\frac{\alpha}{2}, \dots}$ or z_{α} or $t_{\alpha, \dots}$
(or $-z_{\alpha/2}$ or $-t_{\alpha/2, \dots}$) or $-z_{\alpha}$ or $-t_{\alpha}$
2-sided (The one that is relevant) 1-sided

← compare test stat. value to this

P-value = P (test stat. takes a value at least as extreme as its observed value)

↳ compare to α .
All the steps are listed on formula sheet even if test stat. is not. → Review ANOVA steps !!!

Normal Approximation to Binomial

$$X \sim \text{Bin}(n, p)$$

(mean: np
variance: $np(1-p)$)

↓
and $np > 5$, $n(1-p) > 5$ we have:

$$\frac{X - np}{\sqrt{np(1-p)}} \sim N(0, 1).$$

To use this in practice: continuity correction

$$P(X \leq x) = P(X \leq x + 0.5)$$

$$= P\left(Z \leq \frac{(x+0.5) - np}{\sqrt{np(1-p)}}\right)$$

$$P(X = x) = P(X \leq x) - P(X \leq x-1)$$

$$= P\left(Z \leq \frac{(x+0.5) - np}{\sqrt{np(1-p)}}\right) - P\left(Z \leq \frac{(x-0.5) - np}{\sqrt{np(1-p)}}\right)$$

$$P(x_1 \leq X \leq x_2) = P(X \leq x_2) - P(X \leq x_1 - 1)$$

$$= P\left(Z \leq \frac{x_2 + 0.5 - np}{\sqrt{np(1-p)}}\right) - P\left(Z \leq \frac{x_1 - 0.5 - np}{\sqrt{np(1-p)}}\right)$$

$$= P\left(\frac{(x_1 - 0.5) - np}{\sqrt{np(1-p)}} \leq Z \leq \frac{x_2 + 0.5 - np}{\sqrt{np(1-p)}}\right)$$

→ Expand the region when making the continuity correction e.g. if the region would be

$[x_1, x_2]$ as in the last example,

expand to $[x_1 - 0.5, x_2 + 0.5]$ and

then standardize.

e.g. if the region is $\{x\}$, expand to

$[x - 0.5, x + 0.5]$ & standardize.

& e.g. if the region is $[x, \infty)$ or $(-\infty, x]$

expand to $[x-0.5, \infty)$ or $(-\infty, x+0.5]$.
& standardize.

Final Thoughts:

- ① COURSE EVALUATIONS: I'm sure you're very busy, but please take a 10-minute break before the end of Thursday Dec. 6th to review this course — help us to make it better (as well as tell us what we're doing right)!
- ② You've got this — work hard, but keep your energy levels up for the long haul. Practice, practice, practice — but in the end, be confident you'll do your best, and remember that's anyway all anyone can ask of you. **GOOD LUCK !!!**